- Suppose x is a non-negative random variable
- Markov inequality upper bounds the probability that x is large

$$\mathbf{P}\left(x \ge a\right) \le \frac{E[x]}{a} \tag{1}$$

- To use it only information that is needed is the mean of the r.v.
- ► No need for the P.M.F/P.D.F of the r.v.

#### The Chebyshev Inequality

Chebyshev inequality is:

$$\mathbf{P}(|x - E[x]| \ge t) \le \frac{var(x)}{t^2}$$

Only need mean and the variance of the r.v.

- ▶ Proof: Apply the Markov inequality to the new r.v  $|x E[x]|^2$
- ► The Chebyshev inequality can also be written as:

$$\mathbf{P}(x \in [E[x] - t, E[x] + t]) \ge 1 - \frac{var(x)}{t^2}$$
(2)

- ► The interval [E[x] t, E[x] + t] is called confidence interval
- $1 \frac{var(x)}{t^2}$  is called confidence level

### Examples of Markov/Chebyshev inequality

- ► A post office handles, on average, 10,000 letters a day. What can be said about the probability that it will handle at least 15,000 letters tomorrow.
- A post-office handles 10,000 letters per day with a variance of 2,000 letters. What can be said about the probability that this post office handles between 8,000 and 12,000 letters tomorrow? What about the probability that more than 15,000 letters come in?
- A faulty coin with probability of head as 0.2 is tossed 10 times. What is the bound on the probability that 8 heads show up?

#### Sequence, Limits

- Sequence a<sub>n</sub>, limit number a
- a<sub>n</sub> converges to a

$$\lim_{n \to \infty} a_n = a \tag{3}$$

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- "a<sub>n</sub> eventually gets and stays arbitrarily close to a"
- For every *e* > 0, there exists *n*<sub>0</sub> such that for every *n* > *n*<sub>0</sub>, we have

$$|a_n - a_0| \le \epsilon$$

#### Convergence in Probability

- Sequence of random variables Y<sub>n</sub>
- Converges in probability to a number a
- "Almost all of the PDF of Y<sub>n</sub> eventually gets concentrated close to a"
- For every  $\epsilon > 0$

$$\lim_{n\to\infty} \mathbf{P}\left(|Y_n-a|\right) \geq \epsilon = 0$$

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• Example: Does  $Y_n$  converge?  $\mathbf{P}(Y_n = 0) = 1 - \frac{1}{n}$   $\mathbf{P}(Y_n = 1) = \frac{1}{n}$ 

#### Convergence of the sample mean: WLLN

• finite mean  $\mu$  and variance  $\sigma^2$ 

• 
$$M_n = \frac{X_1 + X_2 + \dots + X_n}{n}$$

$$\blacktriangleright E[M_n] =$$

• 
$$var(M_n) =$$

$$\mathbf{P}\left(|M_n - \mu| \ge \epsilon\right) \le \frac{\operatorname{var}(M_n)}{\epsilon^2} = \frac{\sigma^2}{n\,\epsilon^2}$$

#### Polling Problem

- f : fraction of population that ".."
- $i^{th}$  person polled:  $X_i = 1$  if yes,  $X_i = 0$  if no.
- $M_n$  is thus the fraction of yes in our sample.
- Goal: 95% confidence and less than 1% error

$$P(|M_n - f| \ge 0.01) \le 0.05$$

Use Chebyshev inequality:

$$\mathbf{P}(|M_n - f| \ge 0.01) \le \frac{\sigma_x^2}{n \, 0.01^2} \le \frac{1}{4 \, n \, 0.01^2}$$

• n = 50,000. But this is conservative

#### Different Scalings of $M_n$

• 
$$X_1, \ldots, X_n$$
 i.i.d

Three variants of their sum

• 
$$S_n = X_1 + \ldots X_n$$
, variance  $n\sigma^2$ 

• 
$$M_n = \frac{S_n}{n}$$
, variance  $\frac{\sigma^2}{n}$ : WLLN

•  $\frac{S_n}{\sqrt{n}}$ : constant variance  $\sigma^2$ . Asymptotic shape?

#### The Central Limit Theorem

• Standardized 
$$S_n = X_1 + \ldots X_n$$
:

$$Z_n = \frac{S_n - E[S_n]}{\sqrt{n}\sigma}$$

- Zero mean, Unit Variance
- Let Z be a standard normal r.v.
- Theorem: For every c:

$$\mathbf{P}(Z_n \leq c) \to \mathbf{P}(Z \leq c)$$

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# C.L.T

- Treat  $Z_n$ , therefore  $S_n$  as if normal
- universal; only means, variances matter
- accurate computational shortcut
- Can be used even for moderate n
- Used in every walk of life.

#### Polling Problem: Using C.L.T

- f : fraction of population that ".."
- $i^{th}$  person polled:  $X_i = 1$  if yes,  $X_i = 0$  if no.
- $M_n$  is thus the fraction of yes in our sample.
- ▶ Goal: 95% confidence and less than 1% error

$$P(|M_n - f| \ge 0.01) \le 0.05$$

Event of interest:

$$\mathbf{P}\left(|M_n-f|\geq 0.01\right)$$

Using C.L.T

$$\mathbf{P}\left(|M_n - f|\right) > 0.01 \approx \mathbf{P}\left(|Z| > 0.01 \frac{\sqrt{n}}{\sigma}\right)$$

#### Applying C.L.T to Binomal Distribution

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- Let X<sub>i</sub> be Bernoulli(p)
- $S_n = X_1 + X_2 \dots X_n$ , Binomial r.v.
- mean = np, variance = np(1-p)
- ▶ CDF of  $\frac{S_n np}{\sqrt{np(1-p)}}$  → standard normal
- Example: n = 36, p = 0.5, **P** ( $S_n \le 21$ )

$$\sum_{0}^{21} \binom{36}{i} (0.5)^{36} = 0.8785$$

## The $\frac{1}{2}$ correction for binomial approximation

• 
$$P(S_n \le 21) = P(S_n < 22)$$

- ▶ Consider S<sub>n</sub> < 21.5</p>
- De Moivre Laplace CLT (for binomial):
- When the 1/2 correction is used, CLT can also approximate the binomial p.m.f. (not just the binomial CDF)

• 
$$P(S_n = 19) = P(18.5 \le S_n \le 19.5)$$

• 
$$\frac{18.5-18}{3} \le \frac{S_n-18}{3} \le \frac{19.5-18}{3}$$

- $0.17 \le Z_n \le 0.5$
- $P(Z \le 0.5) P(Z \le 0.17) = 0.124.$
- Exact answer: 0.1251